Entropy Rates of a Stochastic Process Best Achievable Data Compression

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Entropy Rates of a Stochastic Process

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- The AEP states that nH(X) bits suffice on the average for n i.i.d. RVs
- What for dependent RVs?
- For stationary processes $H(X_1, X_2, ..., X_n)$ grows (asymptotically) linearly with *n* at a rate $H(\mathcal{X})$ the *entropy rate* of the process
- A stochastic process {X_i}_{i∈I} is an indexed sequence of random variables, X_i : S → X is a RV ∀i ∈ I
- If I ⊆ N, {X₁, X₂,...} is a *discrete stochastic process*, called also a *discrete information source*.
- A discrete stochastic process is characterized by the joint probability mass function

$$P((X_1, X_2, ..., X_n) = (x_1, x_2, ..., x_n)) = p(x_1, x_2, ..., x_n)$$

where $(x_1, x_2, \ldots, x_n) \in \mathcal{X}^n$.

A stochastic process is said to be *stationary* if the joint distribution of any subset of the sequence of random variables is invariant with respect to shifts in the time index

$$P(X_1 = x_1, \dots, X_n = x_n) = P(X_{1+\ell} = x_1, \dots, X_{n+\ell} = x_n)$$
(1)

 $\forall n, \ell \text{ and } \forall x_1, x_2, \dots, x_n \in \mathcal{X}.$

A discrete stochastic process $\{X_1, X_2, ...\}$ is said to be a *Markov chain* or *Markov process* if for n = 1, 2, ...

$$P(X_{n+1} = x_{n+1} | X_n = x_n, X_{n-1} = x_{n-1}, \dots, X_1 = x_1)$$

= $P(X_{n+1} = x_{n+1} | X_n = x_n), \qquad x_1, x_2, \dots, x_n, x_{n+1} \in \mathcal{X}.$ (2)

The joint pmf can be written as

$$p(x_1, x_2, \dots, x_n) = p(x_1)p(x_2|x_1) \dots p(x_n|x_{n-1}).$$
(3)

A Markov chain is said to be *time invariant* (*time homogeneous*) if the conditional probability $p(x_{n+1}|x_n)$ does not depend on *n*; that is for n = 1, 2, ...

$$P(X_{n+1}=b|X_n=a)=P(X_2=b|X_1=a),\quad\forall a,b\in\mathcal{X}.$$
 (4)

This property is assumed unless otherwise stated.

- $\{X_i\}$ Markov chain, X_n is called the *state* at time n
- A time-invariant Markov chain is characterized by its initial state and a *probability transition matrix* $P = [P_{ij}]$, i, j = 1, ..., m, where $P_{ij} = P(X_{n+1} = j | X_n = i)$.
- The Markov chain {*X_n*} is *irreducible* if it is possible to go from any state to another with a probability > 0

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Markov chains IV

- The Markov chain {X_n} is *aperiodic* if ∀ state *a*, the possible times to go from *a* to *a* have highest common factor = 1.
- Markov chains are often described by a directed graph where the edges are labeled by the probability of going from one state to another.
- $p(x_n)$ pmf of the random variable at time n

$$p(x_{n+1}) = \sum_{x_n} p(x_n) P_{x_n x_{n+1}}$$
(5)

• A distribution on the states such that the distribution at time n + 1 is the same as the distribution at time n is called a *stationary distribution* - so called because if the initial state of a Markov chain is drawn according to a stationary distribution, the Markov chain form a stationary process.

• If the finite-state Markov chain is irreducible and periodic, the stationary distribution is unique, and from any starting distribution, the distribution of X_n tends to a stationary distribution as $n \to \infty$.

Example 4

Consider a two-state Markov chain with a probability transition matrix

$$P = \left[egin{array}{cc} 1-lpha & lpha \ eta & 1-eta \end{array}
ight]$$

(Figure 1)

Markov chains VI

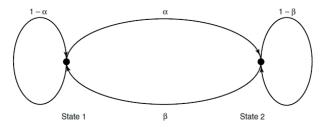


Figure : Two-state Markov chain

The stationary probability is the solution of $\mu P = \mu$ or $(I - P^T)\mu^T = 0$. We add the condition $\mu_1 + \mu_2 = 0$. The solution is

$$\mu_1 = \frac{\beta}{\alpha + \beta}, \quad \mu_2 = \frac{\alpha}{\alpha + \beta}.$$

Click here for a Maple solution Markovex1.html. The entropy of X_n is

$$H(X_n) = H\left(\frac{\beta}{\alpha+\beta}, \frac{\alpha}{\alpha+\beta}\right).$$

The *entropy rate* of a stochastic process $\{X_i\}$ is defined by

$$H(\mathcal{X}) = \lim_{n \to \infty} \frac{1}{n} H(X_1, \dots, X_n)$$
(6)

when the limit exists.

Examples

Typewriter - m equally likely output letters; he(she) can produce mⁿ sequences of length n, all of them equally likely.
 H(X₁,...,X_n) = log mⁿ, and the entropy rate is H(X) = log m bits per symbol.

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② X₁, X₂, ... i.i.d. RVs

$$H(\mathcal{X}) = \lim_{n \to \infty} \frac{H(X_1, \dots, X_n)}{n} = \lim_{n \to \infty} \frac{nH(X_1)}{n} = H(X_1).$$

③ X_1, X_2, \ldots independent, but not identically distributed RVs

$$H(X_1,\ldots,X_n)=\sum_{i=1}^n H(X_i)$$

It is possible that $\frac{1}{n} \sum H(x_i)$ does not exists

$$H'(\mathcal{X}) = \lim_{n \to \infty} H(X_n | X_{n-1}, X_{n-2}, \dots, X_1).$$
(7)

 $H(\mathcal{X})$ is entropy per symbol of the *n* RVs; $H'(\mathcal{X})$ is the conditional entropy of the last RV given the past.

For stationary processes both limits exist and are equal.

Lemma 7

For a stationary stochastic process, $H(X_n|X_{n-1},...,X_1)$ is nonincreasing in n and has a limit $H'(\mathcal{X})$.

Proof.

$$\begin{split} H(X_{n+1}|X_1,X_2,\ldots,X_n) &\leq H(X_{n+1}|X_n,\ldots,X_2) & \text{ conditioning} \\ &= H(X_n|X_{n-1},\ldots,X_1). & \text{ stationarity} \end{split}$$

 $(H(X_n|X_{n-1},\ldots,X_1))_n$ is decreasing and nonnegative, so it has a limit $H'(\mathcal{X}).$

Lemma 8 (Cesáro)

If
$$a_n \rightarrow a$$
 and $b_n = \frac{1}{n} \sum_{i=1}^n a_i$ then $b_n \rightarrow a_i$.

Theorem 9

For a stationary stochastic process $H(\mathcal{X})$ (given by (6)) and $H'(\mathcal{X})$ (given by (7)) exist and

$$\mathcal{H}(\mathcal{X}) = \mathcal{H}'(\mathcal{X}).$$
 (8)

Entropy rate VI

Proof.

By the chain rule,

$$\frac{H(X_1,...,X_n)}{n} = \frac{1}{n} \sum_{i=1}^n H(X_i | X_{i-1},...,X_1).$$

But,

$$H(\mathcal{X}) = \lim_{n \to \infty} \frac{H(X_1, \dots, X_n)}{n}$$

= $\lim_{n \to \infty} \frac{1}{n} \sum_{i=1}^n H(X_i | X_{i-1}, \dots, X_1)$
= $\lim_{n \to \infty} H(X_n | X_{n-1}, \dots, X_1)$ (Lemma 8)
= $H'(\mathcal{X})$ (Lemma 7)

• For a stationary Markov chain, the entropy rate is given by

$$H(\mathcal{X}) = H'(\mathcal{X}) = \lim H(X_n | X_{n-1}, \dots, X_1) = \lim H(X_n | X_{n-1}) = H(X_2 | X_1),$$
(9)

where the conditional entropy is calculated using the given stationary distribution.

• The stationary distribution μ is the solution of the equations

$$\mu_j = \sum_i \mu_i P_{ij}, \ \forall j.$$

• Expression of conditional entropy:

Theorem 10

 $\{X_i\}$ stationary Markov chain with stationary distribution μ and transition matrix P. Let $X_1 \sim \mu$. then the entropy rate is

$$H(\mathcal{X}) = -\sum_{i} \sum_{j} \mu_{i} P_{ij} \log P_{ij}.$$
 (10)

Proof.

$$H(\mathcal{X}) = H(X_2|X_1) = \sum_i \mu_i \left(-\sum_j P_{ij} \log P_{ij}\right).$$

Example 11 (Two-state Markov chain)

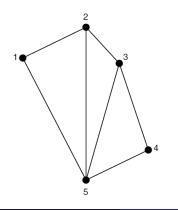
The entropy rate of the two-state Markov chain in Figure 1 is

$$H(\mathcal{X}) = H(X_2|X_1) = \frac{\beta}{\alpha+\beta}H(\alpha) + \frac{\alpha}{\alpha+\beta}H(\beta).$$

Remark. If the Markov chain is irreducible and aperiodic, it has a unique stationary distribution on the states, and any initial distribution tends to the stationary distribution as $n \to \infty$. In this case, even though the initial distribution is not the stationary distribution, the entropy rate, which is defined in terms of long-term behavior, is $H(\mathcal{X})$, as defined in (9) and (10).

Entropy of a random walk on a weighted graph 1

- Assume irreducible and aperiodic (d(i) = 1 for all i) so unique stationary distribution.
- Graph G = (V, E) with *m* nodes labeled $\{1, 2, ..., m\}$ and edges with weight $W_{ij} > 0$.



 Random walk: start at a node, say *i*, and choose next node with probability proportional to edge weight, i.e., *p_{ij}* as

$${\sf P}_{ij} = rac{{\cal W}_{ij}}{\sum_k {\cal W}_{ik}} = rac{{\cal W}_{ij}}{{\cal W}_i}$$

Entropy of a random walk on a weighted graph 2

- Guess that stationary distribution has probability proportional to w_i.
- If $W = \sum_{i,j;j>i} W_{ij}$ then $\sum_i W_i = 2W$, so guess a stationary μ distribution with $\mu_i = w_i/2w$.
- This is stationary since

$$\forall j \ \mu'_j = \sum_i \mu_i P_{ij} = \sum_i \frac{W_i}{2W} \frac{W_{ij}}{W_i} = \frac{1}{2W} \sum_i W_{ij}$$
$$= \frac{W_j}{W} = \mu_j.$$

• Can swap edges elsewhere (i.e., edges between nodes not including *i*), does not change the stationary condition which is local.

What is entropy of this random walk

$$H(\mathcal{X}) = H(X_2|X_1) = -\sum_i \mu_i \sum_j P_{ij} \log P_{ij}$$
$$= -\sum_i \frac{W_i}{2W} \sum_j \frac{W_{ij}}{W_i} \log \frac{W_{ij}}{W_i} = -\sum_i \sum_j \frac{W_{ij}}{2W} \log \frac{W_{ij}}{W_i}$$
$$= -\sum_i \sum_j \frac{W_{ij}}{2W} \log \frac{W_{ij}}{2W} + \sum_i \sum_j \frac{W_{ij}}{2W} \log \frac{W_i}{2W}$$
$$= H\left(\dots, \frac{W_{ij}}{2W}, \dots\right) - H\left(\dots, \frac{W_i}{2W}, \dots\right)$$

- If all the edges have equal weight, the stationary distribution puts weight $E_i/2E$ on node *i*, where E_i is the number of edges emanating from node *i* and *E* is the total number of edges in the graph.
- In this case, the entropy rate of the random walk is

$$H(X) = \log 2E - H\left(\frac{E_1}{2E}, \frac{E_2}{2E}, \dots, \frac{E_m}{2E}\right)$$

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What is entropy of this random walk

• So, the entropy of the random walk is

H(X) = (overall edge uncertainty)

- (overall node uncertainty in stationary condition)

- Intuition: As node entropy decreases while keeping edge uncertainty constant, the network becomes more concentrated, fewer nodes are hubs, and the hubs that remain are widely connected (since edge entropy is fixed).
- In such case (few well connected hubs), it is likely one will land on such a hub (in a random walk) and then will be faced with a wide variety of choice as to where to go next ⇒ increase in overall uncertainty of the walk.
- If node entropy goes up with edge entropy fixed, then many nodes are hubs all with relatively low connectivity, so hitting them doesn't provide much choice => random walk entropy goes down.

- $X_1, X_2, \ldots, X_n, \ldots$ stationary Markov chain, $Y_i = \phi(X_i), H(\mathcal{Y}) = ?$
- in many cases $Y_1, Y_2, \ldots, Y_n, \ldots$ is not a Markov chain, but it is stationary
- Iower bound

Lemma 12 $H(Y_n | Y_{n-1}, ..., Y_2, X_1) \le H(\mathcal{Y}).$ (11)

Functions of Markov chains II

Proof.

For k = 1, 2, ...

$$H(Y_{n}|Y_{n-1},...,Y_{2},X_{1}) \stackrel{(a)}{=} H(Y_{n}|Y_{n-1},...,Y_{2},Y_{1},X_{1})$$

$$\stackrel{(b)}{=} H(Y_{n}|Y_{n-1},...,Y_{2},Y_{1},X_{1},X_{0},X_{-1},...,X_{-k})$$

$$\stackrel{(c)}{=} H(Y_{n}|Y_{n-1},...,Y_{2},Y_{1},X_{1},X_{0},X_{-1},...,X_{-k})$$

$$\stackrel{(d)}{\leq} H(Y_{n}|Y_{n-1},...,Y_{2},Y_{1},Y_{0},...,Y_{-k})$$

$$\stackrel{(e)}{=} H(Y_{n+k+1}|Y_{n+k},...,Y_{1}),$$

(a) follows from the fact that $Y_1 = \phi(X_1)$, (b) from the Markovity, (c) from $Y_i = \phi(X_i)$, (d) conditioning reduces entropy, (e) stationarity.

Proof - continuation.

Since inequality is true for all k, in the limit

$$H(Y_{n}|Y_{n-1},...,Y_{2},X_{1}) \leq \lim_{k} H(Y_{n+k+1}|Y_{n+k},...,Y_{1})$$

= $H(\mathcal{Y}).$

Lemma 13

$$H(Y_n|Y_{n-1},\ldots,Y_2,X_1) - H(Y_n|Y_{n-1},\ldots,Y_2,Y_1,X_1) \to 0.$$
(12)

Functions of Markov chains IV

Proof.

Expression of interval length:

$$H(Y_n|Y_{n-1},\ldots,Y_2,X_1) - H(Y_n|Y_{n-1},\ldots,Y_2,Y_1,X_1) = I(X_1;Y_n|Y_{n-1},\ldots,Y_1).$$

By properties of mutual information,

$$I(X_1; Y_1, \ldots, Y_n) \leq H(X_1),$$

and $I(X_1; Y_1, \ldots, Y_n)$ increases with *n*. Thus, $\lim I(X_1; Y_1, \ldots, Y_n)$ exists and

$$\lim_{n\to\infty}I(X_1;Y_1,\ldots,Y_n)\leq H(X_1).$$

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Functions of Markov chains V

Proof - continuation.

By the chain rule

$$H(X_{1}) \geq \lim_{n \to \infty} I(X_{1}; Y_{1}, \dots, Y_{n})$$

= $\lim_{n \to \infty} \sum_{i=1}^{n} I(X_{1}; Y_{i} | Y_{i-1}, \dots, Y_{1})$
= $\sum_{i=1}^{\infty} I(X_{1}; Y_{i} | Y_{i-1}, \dots, Y_{1})$

The general term of the series must tend to 0

$$\lim I(X_1; Y_n | Y_{n-1}, \ldots, Y_1) = 0.$$

The last two lemmas imply

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Theorem 14

$$X_1, X_2, \dots, X_n, \dots \text{ stationary Markov chain, } Y_i = \phi(X_i)$$
$$H(Y_n | Y_{n-1}, \dots, Y_1, X_1) \le H(\mathcal{Y}) \le H(Y_n | Y_{n-1}, \dots, Y_1)$$
(13)

and

$$\lim H(Y_n | Y_{n-1}, \dots, Y_1, X_1) = H(\mathcal{Y}) = \lim H(Y_n | Y_{n-1}, \dots, Y_1)$$
 (14)

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- We could consider Y_i to be a stochastic function of X_i
- X₁, X₂,..., X_n,... stationary Markov chain, Y₁, Y₂,..., Y_n,... a new process where Y_i is drawn according to p(y_i|x_i), conditionally independent of all the other X_i, j ≠ i

$$p(x^n, y^n) = p(x_1) \prod_{i=1}^{n-1} p(x_{i+1}|x_i) \prod_{i=1}^n p(y_i|x_i).$$

- Y₁, Y₂,..., Y_n,... is called a hidden Markov model (HMM)
- Applied to speech recognition, handwriting recognition, and so on.
- The same argument as for functions of Markov chain works for HMMs.



Thomas M. Cover, Joy A. Thomas, Elements of Information Theory, 2nd edition, Wiley, 2006.



Robert M. Gray, Entropy and Information Theory, Springer, 2009